



DISCRETE APPLIED MATHEMATICS

Discrete Applied Mathematics 127 (2003) 657-663

www.elsevier.com/locate/dam

Bernoulli numbers and the probability of a birthday surprise *

Boaz Tsaban

Department of Mathematics and Computer Science, Bar-Ilan University, Ramat-Gan 52900, Israel

Received 9 January 2001; received in revised form 20 January 2002; accepted 4 February 2002

Abstract

A birthday surprise is the event that, given k uniformly random samples from a sample space of size n, at least two of them are identical. We show that Bernoulli numbers can be used to derive arbitrarily exact bounds on the probability of a birthday surprise. This result can be used in arbitrary precision calculators, and it can be applied to better understand some questions in communication security and pseudorandom number generation.

© 2003 Elsevier Science B.V. All rights reserved.

Keywords: Birthday paradox; Power sums; Bernoulli numbers; Arbitrary precision calculators; Pseudorandomness

1. Introduction

In this note we address the probability β_n^k that in a sample of k uniformly random elements out of a space of size n there exist at least two identical elements. This problem has a long history and a wide range of applications. The term *birthday surprise* for a collision of (at least) two elements in the sample comes from the case n = 365, where the problem can be stated as follows: Assuming that the birthday of people distributes uniformly over the year, what is the probability that in a class of k students, at least two have the same birthday?

It is clear (and well known) that the expected number of collisions (or birthdays) in a sample of k out of n is

$$\binom{k}{2}\frac{1}{n} = \frac{k(k-1)}{2n}.$$

E-mail address: tsaban@macs.biu.ac.il (B. Tsaban).

URL: http://www.cs.biu.ac.il/~tsaban

0166-218X/03/\$- see front matter © 2003 Elsevier Science B.V. All rights reserved. PII: S0166-218X(02)00396-7

Dedicated to my wife Lea on her birthday.

(Indeed, for each distinct i and j in the range $\{1,\ldots,k\}$, let X_{ij} be the random variable taking the value 1 if samples i and j obtained the same value and 0 otherwise. Then the expected number of collisions is $E(\sum_{i\neq j}X_{ij})=\sum_{i\neq j}E(X_{ij})=\sum_{i\neq j}\frac{1}{n}=\binom{k}{2}\frac{1}{n}$.) Thus, 28 students are enough to make the expected number of common birthdays

Thus, 28 students are enough to make the expected number of common birthdays greater than 1. This seemingly surprising phenomenon has got the name *birthday surprise*, or *birthday paradox*.

In several applications, it is desirable to have exact bounds on the probability of a collision. For example, if some electronic application chooses pseudorandom numbers as passwords for its users, it may be a *bad* surprise if two users get the same password by coincidence. It is this term "by coincidence" that we wish to make precise.

2. Bounding the probability of a birthday surprise

When k and n are relatively small, it is a manner of simple calculation to determine β_n^k . The probability that all samples are distinct is

$$\pi_n^k = \left(1 - \frac{1}{n}\right) \left(1 - \frac{2}{n}\right) \cdots \left(1 - \frac{k-1}{n}\right) \tag{1}$$

and $\beta_n^k = 1 - \pi_n^k$. For example, one can check directly that $\beta_{365}^{23} > \frac{1}{2}$, i.e., in a class of 23 students the probability that two share the same birthday is greater than $\frac{1}{2}$. This is another variant of the *Birthday surprise*. ¹

The calculation becomes problematic when k and n are large, both due to precision problems and computational complexity (in cryptographic applications k may be of the order of trillions, i.e., thousands of billions). This problem can be overcome by considering the logarithm of the product:

$$\ln (\pi_n^k) = \sum_{i=1}^{k-1} \ln \left(1 - \frac{i}{n}\right).$$

Since each *i* is smaller than *n*, we can use the Taylor expansion $\ln(1-x) = -\sum_{m=1}^{\infty} x^m/m$ (|x| < 1) to get that

$$-\ln(\pi_n^k) = \sum_{i=1}^{k-1} \sum_{m=1}^{\infty} \frac{(i/n)^m}{m} = \sum_{m=1}^{\infty} \frac{1}{mn^m} \sum_{i=1}^{k-1} i^m.$$
 (2)

(Changing the order of summation is possible because the sums involve positive coefficients.)

The coefficients $p(k-1,m) := \sum_{i=1}^{k-1} i^m$ (which are often called *sums of powers*, or simply *power sums*) play a key role in our estimation of the birthday probability.

¹ To experience this phenomenon experimentally, the reader is referred to [8].

Efficient calculations of the first few power sums go back to ancient mathematics. ² In particular, we have: $\mathfrak{p}(k,1) = k(k+1)/2$, and $\mathfrak{p}(k,2) = k(k+1)/2 + 1/6$. Higher-order power sums can be found recursively using Bernoulli numbers.

The Bernoulli numbers (which are indexed by superscripts) $1=B^0$, B^1 , B^2 , B^3 , B^4 ,... are defined by the formal equation " $B^n = (B-1)^n$ " for n > 1, where the quotation marks indicate that the involved terms are to be expanded in formal powers of B before interpreting. Thus

•
$$B^2 = B^2 - 2B^1 + 1$$
, whence $B^1 = \frac{1}{2}$,

•
$$B^2 = B^2 - 2B^1 + 1$$
, whence $B^1 = \frac{1}{2}$,
• $B^3 = B^3 - 3B^2 + 3B^1 - 1$, whence $B^2 = \frac{1}{6}$,

etc. We thus get that $B^3 = 0$, $B^4 = -\frac{1}{30}$, $B^5 = 0$, $B^6 = \frac{1}{42}$, $B^7 = 0$, and so on. It follows that for each m,

$$\mathfrak{p}(k,m) = \frac{(k+B)^{m+1} - B^{m+1}}{m+1}$$

(Faulhaber's formula [6]) Thus, the coefficients $\mathfrak{p}(k,m)$ can be efficiently calculated for small values of m. In particular, we get that

•
$$\mathfrak{p}(k,3) = \frac{1}{4}k^4 + \frac{1}{2}k^3 + \frac{1}{4}k^2$$
,

•
$$\mathfrak{p}(k,4) = \frac{1}{5}k^5 + \frac{1}{2}k^4 + \frac{1}{3}k^3 - \frac{1}{30}k$$
,

•
$$\mathfrak{p}(k,5) = \frac{1}{6}k^6 + \frac{1}{2}k^5 + \frac{5}{12}k^4 - \frac{1}{12}k^2$$

•
$$\mathfrak{p}(k,6) = \frac{1}{5}k^7 + \frac{1}{5}k^6 + \frac{1}{5}k^5 - \frac{1}{5}k^3 + \frac{1}{42}k$$
,

etc. In order to show that this is enough, we need to bound the tail of the series in Eq. (2). We will achieve this by effectively bounding the power sums.

Lemma 1. Let k be any natural number, and assume that $f:(0,k)\to\mathbb{R}^+$ is such that f''(x) exists, and is nonnegative for all $x \in (0,k)$. Then

$$\sum_{i=1}^{k} f(i) < \int_{0}^{k} f(x + \frac{1}{2}) \, \mathrm{d}x.$$

Proof. For each interval [i, i+1] (i=0,...,k-1), the tangent to the graph of $f(x+\frac{1}{2})$ at $x = i + \frac{1}{2}$ goes below the graph of $f(x + \frac{1}{2})$. This implies that the area of the added part is greater than that of the uncovered part. \Box

Using Lemma 1, we have that for all m > 1,

$$\sum_{i=1}^{k-1} i^m < \int_0^{k-1} (x + \frac{1}{2})^m \, \mathrm{d}x < \frac{(k - 1/2)^{m+1}}{m+1}.$$

² Archimedes (ca. 287-212 BCE) provided a geometrical derivation of a "formula" for the sum of squares [9].

Thus,

$$\sum_{m=N}^{\infty} \frac{\mathfrak{p}(k-1,m)}{mn^m} < \sum_{m=N}^{\infty} \frac{(k-1/2)^{m+1}}{m(m+1)n^m} < \frac{k-1/2}{N(N+1)} \sum_{m=N}^{\infty} \left(\frac{k-1/2}{n}\right)^m$$

$$= \frac{k-\frac{1}{2}}{N(N+1)} \frac{((k-1/2)/n)^N}{1-(k-1/2)/n}$$

$$= \frac{(k-1/2)^{N+1}}{N(N+1)(1-(k-1/2)/n)n^N}. \quad \Box$$
(3)

We thus have the following.

Theorem 2. Let π_n^k denote the probability that all elements in a sample of k elements out of n are distinct. For a natural number N, define

$$\varepsilon_n^k(N) := \frac{(k-1/2)^{N+1}}{N(N+1)(1-(k-1/2)/n)n^N}.$$

Then

$$\sum_{m=1}^{N-1} \frac{\mathfrak{p}(k-1,m)}{mn^m} < -\ln (\pi_n^k) < \sum_{m=1}^{N-1} \frac{\mathfrak{p}(k-1,m)}{mn^m} + \varepsilon_n^k(N).$$

For example, for N = 2 we get

$$\frac{(k-1)k}{2n} < -\ln(\pi_n^k) < \frac{(k-1)k}{2n} + \frac{(k-1/2)^3}{6n^2(1-(k-1/2)/n)}.$$

We demonstrate the tightness of these bounds with a few concrete examples:

Example 3. Let us bound the probability that in a class of five students there exist two sharing the same birthday. Using Theorem 2 with N=2 we get by simple calculation that $\frac{2}{73}<-\ln(\pi_{365}^5)<\frac{2}{73}+\frac{243}{2105320}$, or numerically, 3 0.0273972 < $-\ln(\pi_{365}^5)<0.0275127$. Thus, $0.0270253<\beta_{365}^5<0.0271377$. Repeating the calculations with N=3 yields $0.0271349<\beta_{365}^5<0.0271356$. N=4 shows that $\beta_{365}^5=0.0271355...$

Example 4. We bound the probability that in a class of 73 students there exist two sharing the same birthday, using $N\!=\!2$: $\frac{36}{5} < -\ln(\pi_{365}^{73}) < \frac{36}{5} + \frac{121945}{255792}$, and numerically we get that $0.9992534 < \beta_{365}^{73} < 0.9995882$. For N=3 we get $0.9995365 < \beta_{365}^{73} < 0.9995631$, and for N=8 we get that $\beta_{365}^{73} = 0.9995608...$

In Theorem 2, $\varepsilon_n^k(N)$ converges to 0 exponentially fast with N. In fact, the upper bound is a very good approximation to the actual probability, as can be seen in the

³ All calculations in this paper were performed using the GNU bc calculator [5], with a scale of 500 digits.

above examples. The reason for this is the effectiveness of the bound in Lemma 1 (see [4] for an analysis of this bound as an approximation).

For $k < \sqrt{n}$, we can bound β_n^k directly: Note that for |x| < 1 and odd M, $\sum_{m=0}^M (-x)^m/m! < e^{-x} < \sum_{m=0}^{M+1} (-x)^m/m!$.

Corollary 5. Let β_n^k denote the probability of a birthday surprise in a sample of k out of n, and let $l_N(k,n)$ and $u_N(k,n)$ be the lower and upper bounds from Theorem 2, respectively. Then for all odd M,

$$-\sum_{m=1}^{M+1} \frac{(-l_N(k,n))^m}{m!} < \beta_n^k < -\sum_{m=1}^M \frac{(-u_N(k,n))^m}{m!}.$$

For example, when M = 1 we get that

$$\frac{(k-1)k}{2n} - \frac{(k-1)^2k^2}{4n^2} < \beta_n^k < \frac{(k-1)k}{2n} + \frac{(k-1/2)^3}{6n^2(1-(k-1/2)/n)}.$$
 (4)

The explicit bounds become more complicated when M > 1, but once the lower and upper bounds in Theorem 2 are computed numerically, bounding β_n^k using Corollary 5 is easy. However, Corollary 5 is not really needed in order to deduce the bounds—these can be calculated directly from the bounds of Theorem 2, e.g., using the exponential function built in calculators.

Remark 6.

- (1) It can be proved directly that in fact $\beta_n^k < (k-1)k/2n$ [2]. However, it is not clear how to extend the direct argument to get tighter bounds in a straightforward manner.
- (2) Our lower bound in Eq. (4) compares favorably with the lower bound (1 1/e)((k-1)k)/2n from [3] when $k \le \sqrt{2n/e}$ (when $k > \sqrt{2n/e}$ we need to take larger values of M to get a better approximation).
- (3) p(k-1,m) is bounded from below by $(k-1)^{m+1}/(m+1)$. This implies a slight improvement on Theorem 2.

3. Some applications

3.1. Arbitrary percision calculators

Arbitrary percision calculators do calculations to any desired level of accuracy. Well-known examples are the bc and GNU bc [5] calculators. Theorem 2 allows calculating β_n^k to any desired level of accuracy (in this case, the parameter N will be determined by the required level of accuracy), and in practical time. An example of such calculation appears below (Example 7).

3.2. Cryptography

The probability of a birthday surprise plays an important role in the security analysis of various cryptographic systems. For this purpose, it is common to use the approximation $\beta_n^k \approx k^2/2n$. However, in *concrete security* analysis it is preferred to have exact bounds rather than estimations (see [1] and references therein).

The second item of Remark 6 implies that security bounds derived using earlier methods are tighter than previously thought. The following example demonstrates the tightness of the bounds of Theorem 2 for these purposes.

Example 7. In [2], $\beta_{2^{128}}^{2^{32}}$ is estimated approximately. Using Theorem 2 with N=2, we get that in fact,

$$2^{-65.000000003359036150250796039103} < \beta_{2^{128}}^{2^{32}} < 2^{-65.0000000003359036150250796039042}.$$

With N=3 we get that $\beta_{2^{128}}^{2^{32}}$ lies between

 $2^{-65.000000000335903615025079603904203942942489665995829764250752}$

and

 $2^{-65.000000000335903615025079603904203942942489665995829764250713}\\$

The remarkable tightness of these bounds is due to the fact that 2^{32} is much smaller than 2^{128} .

Another application of our results is for estimations of the quality of approximations such as $\binom{n}{k} \approx n^k/k!$ (when $k \ll n$):

Fact 8.
$$\binom{n}{k} = (n^k/k!)\pi_n^k$$
.

Thus the quality of this approximation is directly related to the quality of the approximation $\pi^k \approx 1$, which is well understood via Theorem 2.

imation $\pi_n^k \approx 1$, which is well understood via Theorem 2. π_n^k appears in many other natural contexts. For example, assume that a function $f:\{0,\ldots,n-1\}\to\{0,\ldots,n-1\}$ is chosen with uniform probability from the set of all such functions, and fix an element $x\in\{0,\ldots,n-1\}$. Then we have the following immediate observation.

Fact 9. The probability that the orbit of x under f has size exactly k is $\pi_n^k k/n$. The probability that the size of the orbit of x is larger than k is simply π_n^k .

These probabilities play an important role in the theory of iterative pseudorandom number generation (see [10] for a typical example).

4. Final remarks and acknowledgments

For a nice account of power sums see [7]. An accessible presentation and proof of Faulhaber's formula appears in [6]. The author thanks John H. Conway for the nice

introduction to Bernoulli numbers, and Ron Adin for reading this note and detecting some typos.

References

- [1] M.Bellare, Practice-Oriented Provable-Security, Lecture Notes in Computer Science, Vol. 1561, Springer, Berlin, 1999, pp. 1–5.
- [2] M. Bellare, O. Goldreich, H. Krawczyk, Stateless evaluation of pseudorandom functions: security beyond the birthday barrier, in: M. Wiener (Ed.), Advances in Cryptology—CRYPTO 99 Proceedings, Lecture Notes in Computer Science, Vol. 1666, Springer, Berlin, 1999, pp. 270–287.
- [3] M. Bellare, J. Kilian, P. Rogaway, The security of the Cipher Block Chaining message authentication code, J. Comput. System Sci. 61 (2000) 362–399.
- [4] B.L. Burrows, R.F. Talbot, Sums of powers of integers, Amer. Math. Monthly 91 (1986) 394-403.
- [5] GNU BC, ftp://prep.ai.mit.edu/pub/gnu/bc/.
- [6] J.H. Conway, R.K. Guy, The Book of Numbers, Copernicus, Springer, New York, 1996, pp. 106-109.
- [7] A.W.F. Edwards, A quick route to sums of powers, Amer. Math. Monthly 93 (1986) 451-455.
- [8] http://www-stat.stanford.edu/~susan/surprise/.
- [9] V.J. Katz, A History of Mathematics: An Introduction, HarperCollins College Publishers, New York, 1993, p. 106.
- [10] A. Shamir, B. Tsaban, Guaranteeing the diversity of pseudorandom generators, Inform. and Comput. 171 (2001) 350–363.