Lecture 4
Gaussian Mixture Models (GMMs)

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Based on the slides of Michael Picheny, Bhuvana Ramabhadran, and Stanley F. Chen presented at the speech recognition course e6870 given at Colombia University Fall 2012.
Can extract features over time (LPC, MFCC, PLP) that . . .
- Characterize info in speech signal in compact form.
- Every $\sim 10$ ms, process window of samples . . .
- To get $\sim 40$ features.

DTW computes distance between feature vectors . . .
- While accounting for nonlinear time alignment.

Learned basic concepts (e.g., distances, shortest paths) . . .
- That will reappear throughout course.
DTW Revisited

\[ w^* = \arg \min_{w \in \text{vocab}} \text{distance}(A'_{\text{test}}, A'_{w}) \]

- **Training**: collect audio \( A_w \) for each word \( w \) in vocab.
  - Generate features \( A'_w \) (template for \( w \)).
- **Test time**: given audio \( A'_{\text{test}} \), convert to \( A'_{\text{test}} \).
  - For each \( w \), compute distance\( (A'_{\text{test}}, A'_{w}) \) using DTW.
  - Return \( w \) with smallest distance.
What are Pros and Cons of DTW?
Pros

- Easy to implement.
- Lots of freedom — can model arbitrary time warpings.
Cons: It's *Ad Hoc*

- Distance measures completely heuristic.
  - Why Euclidean?
  - Weight all dimensions of feature vector equally?
- Warping paths heuristic.
  - Too much freedom not good for robustness?
  - Allowable local paths hand-derived.
- No guarantees of optimality or convergence.
Cons (cont’d)

- Doesn’t scale well?
  - Run DTW for each template in training data.
  - What if large vocabulary? Lots of templates per word?

- Generalization.
  - Doesn’t support mix and match between templates.
Can We Do Better?

- Key insight 1: Learn as much as possible from data.
  - e.g., distance measure; graph weights; graph structure?

- Key insight 2: Use probabilistic modeling.
  - Use well-described theories and models from . . .
  - Probability, statistics, and computer science . . .
  - Rather than arbitrary heuristics with ill-defined properties.
Next Two Main Topics

- **Gaussian Mixture models** (today) — A probabilistic model of

  - Feature vectors associated with a speech sound.
  - Principled distance between test frame . . .
  - And set of template frames.

- **Hidden Markov models** (next week) — A probabilistic model of

  - Time evolution of feature vectors for a speech sound.
  - Principled generalization of DTW.
Part I

Gaussian Distributions
The Scenario

- Given alignment between training feats $X$, test feats $Y$.
  - Warping functions $\tau_x(t), \tau_y(t), t = 1, \ldots, T$.
  - *i.e.*, time $\tau_x(t)$ in $X$ aligns with time $\tau_y(t)$ in $Y$.
- Total distance is sum of distance between aligned vectors.

$$\text{distance}_{\tau_x, \tau_y}(X, Y) = \sum_{t=1}^{T} \text{framedist}(x_{\tau_x(t)}, y_{\tau_y(t)})$$
The Scenario

- Computing frame distance for pair of frames is easy.
  \[
  \text{framedist}(x_{\tau x}(t), y_{\tau y}(t))
  \]
- Imagine 2d feature vectors instead of 40d for visualization.
What if instead of one training sample, have many?

\[
\text{framedist}((x_{\tau_{x1}(t)}, x_{\tau_{x2}(t)}, x_{\tau_{x3}(t)}, \ldots); y_{\tau_y(t)})
\]
Ideas

- Average training samples; compute Euclidean distance.
- Find best match over all training samples.
- Make *probabilistic model* of training samples.
Probabilistic Modeling

- Old paradigm:
  \[ w^* = \arg \min_{w \in \text{vocab}} \text{distance}(A'_\text{test}, A'_w) \]

- New paradigm:
  \[ w^* = \arg \min_{w \in \text{vocab}} - \log P(A'_\text{test}|w) \]

- \( P(A'|w) \) is (relative) frequency with which \( w \) ... 
- Is realized as feature vector \( A' \).
Why Probabilistic Modeling?

- If can estimate $P(A'|w)$ perfectly . . .
  - Can perform classification optimally!
- e.g., two-class classification (w/ classes equally frequent)
  - Choose yes iff $P(A'_{test}|yes) > P(A'_{test}|no)$.
  - This is best you can do!
Why Probabilistic Modeling?

- In limit of infinite data, . . .
  - Can estimate probabilities perfectly (*consistency*).
- In real world situations (e.g., sparse data) . . .
  - No guarantees.
  - Still, better to follow principles imperfectly . . .
  - Than to not have principles at all.
Where Are We?

1. Gaussians in One Dimension
2. Gaussians in Multiple Dimensions
3. Estimating Gaussians From Data
Estimate $P(x_1, x_2)$, the “frequency” …

That training sample occurs at location $(x_1, x_2)$. 

Problem Formulation, Two Dimensions
Let’s Start With One Dimension

- Estimate $P(x)$, the “frequency” …
- That training sample occurs at location $x$. 
The Gaussian or Normal Distribution

\[ P_{\mu,\sigma^2}(x) = \mathcal{N}(\mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \]

- Parametric distribution with two parameters:
  - \( \mu = \) mean (the center of the data).
  - \( \sigma^2 = \) variance (how wide data is spread).
Visualization

- **Density function:**

![Density function graph]

- **Sample from distribution:**

![Sample from distribution graph]
Properties of Gaussian Distributions

- Is valid distribution.

\[
\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \, dx = 1
\]

- Central Limit Theorem: Sums of large numbers of identically distributed random variables tend to Gaussian.
  - Lots of different types of data look “bell-shaped”.
- Sums and differences of Gaussian random variables . . .
  - Are Gaussian.
- If \( X \) is distributed as \( \mathcal{N}(\mu, \sigma^2) \) . . .
  - \( aX + b \) is distributed as \( \mathcal{N}(a\mu + b, (a\sigma)^2) \).
- Negative log looks like weighted Euclidean distance!

\[
\ln \sqrt{2\pi\sigma} + \frac{(x - \mu)^2}{2\sigma^2}
\]
Where Are We?

1. Gaussians in One Dimension
2. Gaussians in Multiple Dimensions
3. Estimating Gaussians From Data
Gaussians in Two Dimensions

\[ \mathcal{N}(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1 - r^2}} e^{-\frac{1}{2(1-r^2)} \left( \frac{(x_1 - \mu_1)^2}{\sigma_1^2} - \frac{2x_1x_2}{\sigma_1\sigma_2} + \frac{(x_2 - \mu_2)^2}{\sigma_2^2} \right)} \]

- If \( r = 0 \), simplifies to

\[ \frac{1}{\sqrt{2\pi\sigma_1}} e^{-\frac{(x_1 - \mu_1)^2}{2\sigma_1^2}} \quad \frac{1}{\sqrt{2\pi\sigma_2}} e^{-\frac{(x_2 - \mu_2)^2}{2\sigma_2^2}} = \mathcal{N}(\mu_1, \sigma_1^2)\mathcal{N}(\mu_2, \sigma_2^2) \]

- \text{i.e., like generating each dimension independently.}
Example: $r = 0, \sigma_1 = \sigma_2$

- $x_1, x_2$ uncorrelated.
- Knowing $x_1$ tells you nothing about $x_2$. 
Example: $r = 0$, $\sigma_1 \neq \sigma_2$

- $x_1, x_2$ can be uncorrelated and have unequal variance.
Example: $r > 0, \sigma_1 \neq \sigma_2$

- $x_1, x_2$ correlated.
- Knowing $x_1$ tells you something about $x_2$. 

![Scatter plot of correlated data points](image-url)
If we write following matrix:

\[
\Sigma = \begin{pmatrix}
\sigma_1^2 & r\sigma_1\sigma_2 \\
r\sigma_1\sigma_2 & \sigma_2^2
\end{pmatrix}
\]

then another way to write two-dimensional Gaussian is:

\[
\mathcal{N}(\mu, \Sigma) = \frac{1}{(2\pi)^{d/2}|\Sigma|^{1/2}} \ e^{-\frac{1}{2}(x-\mu)^T\Sigma^{-1}(x-\mu)}
\]

where \( x = (x_1, x_2) \), \( \mu = (\mu_1, \mu_2) \).

More generally, \( \mu \) and \( \Sigma \) can have arbitrary numbers of components.

**Multivariate** Gaussians.
Diagonal and Full Covariance Gaussians

- Let’s say have 40d feature vector.
  - How many parameters in covariance matrix $\Sigma$?
  - The more parameters, . . .
  - The more data you need to estimate them.
- In ASR, usually assume $\Sigma$ is diagonal $\Rightarrow d$ params.
  - This is why like having uncorrelated features!
- (Research direction: is there something in between?)
Computing Gaussian Log Likelihoods

- Why \( \log \) likelihoods?
- Full covariance:

\[
\log P(x) = -\frac{d}{2} \ln(2\pi) - \frac{1}{2} \ln |\Sigma| - \frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu)
\]

- Diagonal covariance:

\[
\log P(x) = -\frac{d}{2} \ln(2\pi) - \sum_{i=1}^{d} \ln \sigma_i - \frac{1}{2} \sum_{i=1}^{d} \left( x_i - \mu_i \right)^2 / \sigma_i^2
\]

- Again, note similarity to weighted Euclidean distance.
- Terms on left independent of \( x \); precompute.
- A few multiplies/adds per dimension.
Where Are We?

1. Gaussians in One Dimension
2. Gaussians in Multiple Dimensions
3. Estimating Gaussians From Data
Estimating Gaussians

- Give training data, how to choose parameters $\mu$, $\Sigma$?
- Find parameters so that resulting distribution …
  - “Matches” data as well as possible.
- Sample data: height, weight of baseball players.
Maximum-Likelihood Estimation (Univariate)

- One criterion: data “matches” distribution well . . .
  - If distribution assigns high likelihood to data.
- Likelihood of string of observations $x_1, x_2, \ldots, x_N$ is . . .
  - Product of individual likelihoods.

$$L(x_1^N | \mu, \sigma) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}}$$

- Maximum likelihood estimation: choose $\mu, \sigma$ . . .
  - That maximizes likelihood of training data.

$$(\mu, \sigma)_{\text{MLE}} = \arg \max_{\mu, \sigma} L(x_1^N | \mu, \sigma)$$
Why Maximum-Likelihood Estimation?

- Assume we have “correct” model form.
- Then, in presence of infinite training samples . . .
  - ML estimates approach “true” parameter values.
  - For most models, MLE is asymptotically consistent, unbiased, and efficient.
- ML estimation is easy for many types of models.
  - Count and normalize!
What is ML Estimate for Gaussians?

- Much easier to work with log likelihood \( L = \ln \mathcal{L} \):

\[
L(x_1^N | \mu, \sigma) = -\frac{N}{2} \ln 2\pi\sigma^2 - \frac{1}{2} \sum_{i=1}^{N} \frac{(x_i - \mu)^2}{\sigma^2}
\]

- Take partial derivatives w.r.t. \( \mu, \sigma \):

\[
\frac{\partial L(x_1^N | \mu, \sigma)}{\partial \mu} = \sum_{i=1}^{N} \frac{(x_i - \mu)}{\sigma^2} \\
\frac{\partial L(x_1^N | \mu, \sigma)}{\partial \sigma^2} = -\frac{N}{2\sigma^2} + \sum_{i=1}^{N} \frac{(x_i - \mu)^2}{\sigma^4}
\]

- Set equal to zero; solve for \( \mu, \sigma^2 \).

\[
\mu = \frac{1}{N} \sum_{i=1}^{N} x_i \\
\sigma^2 = \frac{1}{N} \sum_{i=1}^{N} (x_i - \mu)^2
\]
What is ML Estimate for Gaussians?

- Multivariate case.

\[ \mu = \frac{1}{N} \sum_{i=1}^{N} x_i \]

\[ \Sigma = \frac{1}{N} \sum_{i=1}^{N} (x_i - \mu)^T (x_i - \mu) \]

- What if diagonal covariance?
  - Estimate params for each dimension independently.
Example: ML Estimation

- Heights (in.) and weights (lb.) of 1033 pro baseball players.
- Noise added to hide discretization effects.
- \(\sim stanchen/e6870/data/mlb_data.dat\)

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<td>...</td>
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</table>
Example: ML Estimation
Example: Diagonal Covariance

\[ \mu_1 = \frac{1}{1033} (74.34 + 73.92 + 72.01 + \cdots) = 73.71 \]

\[ \mu_2 = \frac{1}{1033} (181.29 + 213.79 + 209.52 + \cdots) = 201.69 \]

\[ \sigma_1^2 = \frac{1}{1033} \left[ (74.34 - 73.71)^2 + (73.92 - 73.71)^2 + \cdots \right] \]

\[ = 5.43 \]

\[ \sigma_2^2 = \frac{1}{1033} \left[ (181.29 - 201.69)^2 + (213.79 - 201.69)^2 + \cdots \right] \]

\[ = 440.62 \]
Example: Diagonal Covariance
Example: Full Covariance

- Mean; diagonal elements of covariance matrix the same.

\[ \Sigma_{12} = \Sigma_{21} = \frac{1}{1033} \left[ (74.34 - 73.71) \times (181.29 - 201.69) + \\
(73.92 - 73.71) \times (213.79 - 201.69) + \cdots \right] \]

\[ = 25.43 \]

\[ \mu = [73.71 \hspace{1em} 201.69] \]

\[ \Sigma = \begin{bmatrix}
5.43 & 25.43 \\
25.43 & 440.62
\end{bmatrix} \]
Example: Full Covariance
Recap: Gaussians

- Lots of data “looks” Gaussian.
  - Central limit theorem.
- ML estimation of Gaussians is easy.
  - Count and normalize.
- In ASR, mostly use diagonal covariance Gaussians.
  - Full covariance matrices have too many parameters.
Part II

Gaussian Mixture Models
Problems with Gaussian Assumption
Problems with Gaussian Assumption

- Sample from MLE Gaussian trained on data on last slide.
- Not all data is Gaussian!
Problems with Gaussian Assumption

What can we do? What about two Gaussians?

\[ P(x) = p_1 \times \mathcal{N}(\mu_1, \Sigma_1) + p_2 \times \mathcal{N}(\mu_2, \Sigma_2) \]

where \( p_1 + p_2 = 1 \).
Gaussian Mixture Models (GMM’s)

More generally, can use arbitrary number of Gaussians:

\[
P(x) = \sum_j p_j \frac{1}{(2\pi)^{d/2}|\Sigma_j|^{1/2}} \ e^{-\frac{1}{2}(x-\mu_j)^T\Sigma_j^{-1}(x-\mu_j)}
\]

where \(\sum_j p_j = 1\) and all \(p_j \geq 0\).

Also called *mixture* of Gaussians.

Can approximate any distribution of interest pretty well . . .

If just use enough component Gaussians.
Example: Some Real Acoustic Data
Example: 10-component GMM (Sample)
Example: 10-component GMM ($\mu$’s, $\sigma$’s)
Given training data, how to estimate parameters . . .
  * i.e., the $\mu_j$, $\Sigma_j$, and mixture weights $p_j$ . . .
  * To maximize likelihood of data?

No closed-form solution.
  * Can’t just count and normalize.

Instead, must use an optimization technique . . .
  * To find good local optimum in likelihood.
  * Gradient search
  * Newton’s method

Tool of choice: The Expectation-Maximization Algorithm.
1. The Expectation-Maximization Algorithm

2. Applying the EM Algorithm to GMM’s
This is another key thing to remember from course.

Used to train GMM’s, HMM’s, and lots of other things.

Key paper in 1977 by Dempster, Laird, and Rubin [2].
What Does The EM Algorithm Do?

- Finds ML parameter estimates for models . . .
  - With *hidden* variables.
- Iterative hill-climbing method.
  - Adjusts parameter estimates in each iteration . . .
  - Such that likelihood of data . . .
  - Increases (weakly) with each iteration.
- Actually, finds local optimum for parameters in likelihood.
What is a Hidden Variable?

- A random variable that isn’t observed.
- Example: in GMMs, output prob depends on . . .
  - The mixture component that generated the observation
  - But you can’t observe it
- So, to compute prob of observed $x$, need to sum over . . .
  - All possible values of hidden variable $h$:

$$ P(x) = \sum_h P(h, x) = \sum_h P(h) P(x|h) $$
Mixtures and Hidden Variables

Consider probability that is mixture of probs, e.g., a GMM:

\[ P(x) = \sum_j p_j \mathcal{N}(\mu_j, \Sigma_j) \]

Can be viewed as hidden model.

- \( h \leftrightarrow \) Which component generated sample.
- \( P(h) = p_j; \ P(x|h) = \mathcal{N}(\mu_j, \Sigma_j). \)

\[ P(x) = \sum_h P(h)P(x|h) \]
The Basic Idea

- If nail down “hidden” value for each $x_i, \ldots$
  - Model is no longer hidden!
  - *e.g.*, data partitioned among GMM components.
- So for each data point $x_i$, assign single hidden value $h_i$.
  - Take $h_i = \text{arg max}_h P(h) P(x_i|h)$.
  - *e.g.*, identify GMM component generating each point.
- Easy to train parameters in non-hidden models.
  - Update parameters in $P(h), P(x|h)$.
  - *e.g.*, count and normalize to get MLE for $\mu_j, \Sigma_j, \rho_j$.
- Repeat!
The Basic Idea

- **Hard decision:**
  - For each $x_i$, assign single $h_i = \arg\max_h P(h, x_i)$ ...
  - With count 1.

- **Soft decision:**
  - For each $x_i$, compute for every $h$ ...
  - the Posterior prob $\tilde{P}(h|x_i) = \frac{P(h,x_i)}{\sum_h P(h,x_i)}$.
  - Also called the “fractional count”
  - *e.g.*, partition event across *every* GMM component.

- Rest of algorithm unchanged.
The Basic Idea

- Initialize parameter values somehow.
- For each iteration . . .
- Expectation step: compute posterior (count) of $h$ for each $x_i$. 
  \[
  \tilde{P}(h|x_i) = \frac{P(h, x_i)}{\sum_h P(h, x_i)}
  \]
- Maximization step: update parameters.
  - Instead of data $x_i$ with hidden $h$, pretend . . .
  - *Non-hidden* data where . . .
  - (Fractional) count of each $(h, x_i)$ is $\tilde{P}(h|x_i)$. 

Example: Training a 2-component GMM

- Two-component univariate GMM; 10 data points.
- The data: $x_1, \ldots, x_{10}$

$$8.4, 7.6, 4.2, 2.6, 5.1, 4.0, 7.8, 3.0, 4.8, 5.8$$

- Initial parameter values:

<table>
<thead>
<tr>
<th>$p_1$</th>
<th>$\mu_1$</th>
<th>$\sigma_1^2$</th>
<th>$p_2$</th>
<th>$\mu_2$</th>
<th>$\sigma_2^2$</th>
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</tbody>
</table>

- Training data; densities of initial Gaussians.
### The E Step

| $x_i$ | $p_1 \cdot \mathcal{N}_1$ | $p_2 \cdot \mathcal{N}_2$ | $P(x_i)$ | $\tilde{P}(1|x_i)$ | $\tilde{P}(2|x_i)$ |
|-------|--------------------------|--------------------------|----------|------------------|------------------|
| 8.4   | 0.0000                   | 0.0749                   | 0.0749   | 0.000            | 1.000            |
| 7.6   | 0.0003                   | 0.1666                   | 0.1669   | 0.002            | 0.998            |
| 4.2   | 0.1955                   | 0.0040                   | 0.1995   | 0.980            | 0.020            |
| 2.6   | 0.0749                   | 0.0000                   | 0.0749   | 1.000            | 0.000            |
| 5.1   | 0.1089                   | 0.0328                   | 0.1417   | 0.769            | 0.231            |
| 4.0   | 0.1995                   | 0.0022                   | 0.2017   | 0.989            | 0.011            |
| 7.8   | 0.0001                   | 0.1448                   | 0.1450   | 0.001            | 0.999            |
| 3.0   | 0.1210                   | 0.0001                   | 0.1211   | 0.999            | 0.001            |
| 4.8   | 0.1448                   | 0.0177                   | 0.1626   | 0.891            | 0.109            |
| 5.8   | 0.0395                   | 0.0971                   | 0.1366   | 0.289            | 0.711            |

\[
\tilde{P}(h|x_i) = \frac{P(h, x_i)}{\sum_h P(h, x_i)} = \frac{p_h \cdot \mathcal{N}_h}{P(x_i)} \quad h \in \{1, 2\}
\]
View: have *non-hidden* corpus for each component GMM.

For $h$th component, have $\tilde{P}(h|x_i)$ counts for event $x_i$.

Estimating $\mu$: fractional events.

$$\mu = \frac{1}{N} \sum_{i=1}^{N} x_i \quad \Rightarrow \quad \mu_h = \frac{1}{\sum_i \tilde{P}(h|x_i)} \sum_{i=1}^{N} \tilde{P}(h|x_i) x_i$$

$$\mu_1 = \frac{1}{0.000 + 0.002 + 0.980 + \cdots \times (0.000 \times 8.4 + 0.002 \times 7.6 + 0.980 \times 4.2 + \cdots)}$$

$$= 3.98$$

Similarly, can estimate $\sigma_h^2$ with fractional events.
What about the mixture weights $p_h$?

To find MLE, count and normalize!

$$p_1 = \frac{0.000 + 0.002 + 0.980 + \cdots}{10} = 0.59$$
The End Result

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<th>$\sigma_1^2$</th>
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First Few Iterations of EM

iter 0

iter 1

iter 2
Later Iterations of EM

iter 2

iter 3

iter 10
Why the EM Algorithm Works [3]

- \( \mathbf{x} = (x_1, x_2, \ldots) = \) whole training set; \( \mathbf{h} = \) hidden.
- \( \theta = \) parameters of model.
- Objective function for MLE: (log) likelihood.

\[
L(\theta) = \log P(\mathbf{x}|\theta) = \log \sum_{\mathbf{h}} P(\mathbf{h}, \mathbf{x}|\theta)
\]

- Alternate objective function
  - Will show maximizing this equivalent to above
    \[
    F(\tilde{P}, \theta) = L(\theta) - D(\tilde{P} \parallel P_\theta)
    \]
  - \( P_\theta(\mathbf{h}|\mathbf{x}) = \) posterior over hidden.
  - \( \tilde{P}(\mathbf{h}) = \) distribution over hidden to be optimized . . .
  - \( D(\cdot \parallel \cdot) = \) Kullback-Leibler divergence.
Why the EM Algorithm Works

\[ F(\tilde{P}, \theta) = L(\theta) - D(\tilde{P} \parallel P_\theta) \]

Outline of proof:
- Show that both E step and M step improve \( F(\tilde{P}, \theta) \).
- Will follow that likelihood \( L(\theta) \) improves as well.
The E Step

\[ F(\tilde{P}, \theta) = L(\theta) - D(\tilde{P} \parallel P_\theta) \]

- Properties of KL divergence.
  - Nonnegative; and zero iff \( \tilde{P} = P_\theta \).
- What is best choice for \( \tilde{P}(h) \)?
  - Compute the current posterior \( P_\theta(h|x) \).
  - Set \( \tilde{P}(h) \) equal to this posterior.
- Since \( L(\theta) \) is not function of \( \tilde{P} \) . . .
  - \( F(\tilde{P}, \theta) \) can only improve in E step.
The M Step

Lemma:

\[ F(\tilde{P}, \theta) = E_{\tilde{P}}[\log P(h, x|\theta)] + H(\tilde{P}) \]

Proof:

\[ F(\tilde{P}, \theta) = L(\theta) - D(\tilde{P} \parallel P_\theta) \]

\[ = \log P(x|\theta) - \sum_h \tilde{P}(h) \log \frac{\tilde{P}(h)}{P(h|x, \theta)} \]

\[ = \log P(x|\theta) - \sum_h \tilde{P}(h) \log \frac{\tilde{P}(h)P(x|\theta)}{P(h, x|\theta)} \]

\[ = \sum_h \tilde{P}(h) \log P(h, x|\theta) - \sum_h \tilde{P}(h) \log \tilde{P}(h) \]

\[ = E_{\tilde{P}}[\log P(h, x|\theta)] + H(\tilde{P}) \]
The M Step (cont’d)

\[ F(\tilde{P}, \theta) = E_{\tilde{P}}[\log P(h, x|\theta)] + H(\tilde{P}) \]

- \( E_{\tilde{P}}[\cdots] = \) log likelihood of non-hidden corpus . . .
  - Where each \( h \) gets \( \tilde{P}(h) \) counts.
- \( H(\tilde{P}) = \text{entropy of distribution} \tilde{P}(h). \)

What do we do in M step?

- Pick \( \theta \) to maximize term on left
- Note this is just MLE of non-hidden corpus . . .
- Since we chose an estimate for \( h \) from the E step.

- Since \( H(\tilde{P}) \) is not function of \( \theta \) . . .
- \( F(\tilde{P}, \theta) \) can only improve in M step.
Why the EM Algorithm Works

- Observation: $F(\tilde{P}, \theta) = L(\theta)$ after E step (set $\tilde{P} = P_\theta$).

$$F(\tilde{P}, \theta) = L(\theta) - D(\tilde{P} \parallel P_\theta)$$

- If $F(\tilde{P}, \theta)$ improves with each iteration . . .
  - And $F(\tilde{P}, \theta) = L(\theta)$ after each E step . . .
  - $L(\theta)$ improves after each iteration.

- There you go!
Discussion

- EM algorithm is elegant and general way to ...  
  - Train parameters in hidden models ...  
  - To optimize likelihood.

- Only finds local optimum.
  - *Seeding* is of paramount importance.

- Generalized EM algorithm.
  - $F(\tilde{P}, \theta)$ just needs to improve *some* in each step.  
  - *i.e.*, $\tilde{P}(h)$ in E step need not be exact posterior.  
  - *i.e.*, $\theta$ in M step need not be ML estimate.  
  - *e.g.*, can optimize *Viterbi* likelihood.
Where Are We?

1. The Expectation-Maximization Algorithm

2. Applying the EM Algorithm to GMM’s
Another Example Data Set
Question: How Many Gaussians?

- Method 1 (most common): Guess!
- Method 2: Bayesian Information Criterion (BIC)[1].
  - Penalize likelihood by number of parameters.

\[
BIC(C_k) = \sum_{j=1}^{k} \left\{ -\frac{1}{2} n_j \log |\Sigma_j| \right\} - Nk \left( d + \frac{1}{2} d(d + 1) \right)
\]

- \( k \) = Gaussian components.
- \( d \) = dimension of feature vector.
- \( n_j \) = data points for Gaussian \( j \); \( N \) = total data points.
- View GMM as way of coding data for transmission.
  - Cost of transmitting model $\Leftrightarrow$ number of params.
  - Cost of transmitting data $\Leftrightarrow$ log likelihood of data.
- Choose number of Gaussians to minimize cost.

Figure 1. Different degrees of complexity in phone ER-3 and KD-3
Question: How To Initialize Parameters?

- Set mixture weights $p_j$ to $1/k$ (for $k$ Gaussians).
- Pick N data points at random and . . .
  - Use them to seed initial values of $\mu_j$.
- Set all $\sigma$’s to arbitrary value . . .
  - Or to global variance of data.
- Extension: generate multiple starting points.
  - Pick one with highest likelihood.
Another Way: Splitting

- Start with single Gaussian, MLE.
- Repeat until hit desired number of Gaussians:
  - Double number of Gaussians by perturbing means . . .
  - Of existing Gaussians by $\pm \epsilon$.
  - Run several iterations of EM.
Question: How Long To Train?

- \textit{i.e.}, how many iterations of EM?
- Guess.

Look at performance on training data.
- Stop when change in log likelihood per event ... 
- Is below fixed threshold.

Look at performance on held-out data.
- Stop when performance no longer improves.
The Data Set
Sample From Best 1-Component GMM
The Data Set, Again
20-Component GMM Trained on Data
20-Component GMM $\mu$'s, $\sigma$'s
Acoustic Feature Data Set
5-Component GMM; Starting Point A
5-Component GMM; Starting Point B
5-Component GMM; Starting Point C
Solutions With Infinite Likelihood

- Consider log likelihood; two-component 1d Gaussian.

\[
\sum_{i=1}^{N} \ln \left( p_1 \frac{1}{\sqrt{2\pi\sigma_1}} e^{-\frac{(x_i - \mu_1)^2}{2\sigma_1^2}} + p_2 \frac{1}{\sqrt{2\pi\sigma_2}} e^{-\frac{(x_i - \mu_2)^2}{2\sigma_2^2}} \right)
\]

- If \( \mu_1 = x_1 \), above reduces to

\[
\ln \left( \frac{1}{2\sqrt{2\pi\sigma_1}} + \frac{1}{2\sqrt{2\pi\sigma_2}} e^{\frac{1}{2} \frac{(x_1 - \mu_2)^2}{\sigma_2^2}} \right) + \sum_{i=2}^{N} \ldots
\]

which goes to \( \infty \) as \( \sigma_1 \to 0 \).

- Only consider finite local maxima of likelihood function.
  - Variance flooring.
  - Throw away Gaussians with “count” below threshold.
GMM’s are effective for modeling arbitrary distributions.
- State-of-the-art in ASR for decades.
- The EM algorithm is primary tool for training GMM’s.
  - Very sensitive to starting point.
  - Initializing GMM’s is an art.


Given test sample, find nearest training sample.

\[ w^* = \arg \min_{w \in \text{vocab}} \text{distance}(A'_{\text{test}}, A'_w) \]

- Total distance between training and test sample . . .
  - Is sum of distances between aligned frames.
    \[ \text{distance}_{\tau_x, \tau_y}(X, Y) = \sum_{t=1}^{T} \text{framedist}(x_{\tau_x(t)}, y_{\tau_y(t)}) \]

- Goal: move from \textit{ad hoc} distances to probabilities.
Assume many training templates for each word.

- Calc distance between set of training frames . . .
- And test frame.

\[
\text{framedist}((x_1, x_2, \ldots, x_D); y)
\]

Idea: use \(x_1, x_2, \ldots, x_D\) to train GMM: \(P(x)\).

\[
\text{framedist}((x_1, x_2, \ldots, x_D); y) \Rightarrow - \log P(y)
\]
What’s Next: Hidden Markov Models

- Replace DTW with probabilistic counterpart.
- Together, GMM’s and HMM’s comprise . . .
  - Unified probabilistic framework.
- Old paradigm:
  \[ w^* = \arg \min_{w \in \text{vocab}} \text{distance}(A'_{test}, A'_w) \]
- New paradigm:
  \[ w^* = \arg \max_{w \in \text{vocab}} P(A'_{test}|w) \]